

WORKSHEET 4

In all the exercises,  $(\Omega, \mathcal{A}, \mathbb{P})$  denotes the current probability space and  $(B_t)_{t \geq 0}$  a (real) Brownian motion.

1. CONTINUOUS MARTINGALES

**Exercise 1.** Let  $X$  be a r.v. with finite expectation and  $(\mathcal{F}_t)_{t \geq 0}$  be a filtration. Prove that  $(\mathbb{E}[X|\mathcal{F}_t])_{t \geq 0}$  is a  $(\mathcal{F}_t)_{t \geq 0}$ -martingale.

**Exercise 2.** *Process with independent increments.* Let  $(X_t)_{t \geq 0}$  be an integrable process with independent increments, that is for any  $t > s$ ,  $X_t - X_s$  is independent of  $\sigma(X_u, u \leq s)$ , such that for any  $t \geq 0$ ,  $\mathbb{E}[X_t] = \mathbb{E}[X_0]$ . Prove that:

- (1)  $(X_t)_{t \geq 0}$  is a martingale,
- (2) if for any  $t \geq 0$ ,  $\mathbb{E}[X_t^2] < \infty$  then  $(X_t^2 - \mathbb{E}[X_t^2])_{t \geq 0}$  is a martingale,
- (3) if, for some  $\lambda \in \mathbb{R}$  and for any  $t \geq 0$ ,  $\mathbb{E}[e^{\lambda X_t}] < \infty$  then  $(Z_t)_{t \geq 0}$  is a martingale, where

$$Z_t = e^{\lambda X_t} / \mathbb{E} [ e^{\lambda X_t} ] .$$

**Exercise 3.** Let  $M$  be a continuous non negative martingale such that  $M_0 = a > 0$  and  $\lim_{t \rightarrow \infty} M_t = 0$  almost surely.

- (1) For  $y > 0$ , let  $T_y = \inf\{t \geq 0, M_t = y\}$ . Prove that  $\mathbb{P}(T_y < \infty) = a/y$ .
- (2) Prove that  $\sup_{t \geq 0} M_t \sim \frac{a}{U}$  where  $U \sim \mathcal{U}([0, 1])$ .

2. MARTINGALE AND BROWNIAN MOTION

**Exercise 4.** Let  $B_1$  and  $B_2$  be two independent Brownian motions. Prove that the process  $X = B_1 B_2$  is a martingale with respect to the filtration  $(\mathcal{F}_t^{(B_1, B_2)})_{t \geq 0}$ .

**Exercise 5.** Let  $B$  be a Brownian motion started in  $x \in \mathbb{R}$ . Let  $a \leq x \leq b$  and define  $T_a = \inf\{t \geq 0, B_t = a\}$  and  $T_b = \inf\{t \geq 0, B_t = b\}$ .

- (1) Prove that  $B$  is a martingale. Is  $B$  uniformly integrable?
- (2) Using Doob's optional stopping theorem, prove that

$$\mathbb{P}(T_b < T_a) = \frac{x - a}{b - a} .$$

- (3) We suppose now that  $x = 0$ .
  - (a) Prove that the process  $(B_t^2 - t)_{t \geq 0}$  is a  $(\mathcal{F}_t^B)_{t \geq 0}$ -martingale.
  - (b) Use the previous results to show that  $\mathbb{E}[T_a \wedge T_b] = |a|b$ . What can you say about  $\mathbb{E}[T_a]$ ?

**Exercise 6.** Let  $I = -\inf_{0 \leq t \leq T_1} B_t$  where  $T_1 = \inf\{t \geq 0, B_t = 1\}$ . Prove that  $I$  is a continuous r.v. with density  $f(x) = \frac{1}{(1+x)^2} \mathbf{1}_{x \geq 0}$ .

**Exercise 7.** Let  $B$  be a Brownian Motion started in 0 and  $\lambda$  be a real number and define  $M_t^\lambda = e^{\lambda B_t - \lambda^2 t/2}$ .



(1) Prove that the process  $(M_t^\lambda)_{t \geq 0}$  is a  $(\mathcal{F}_t^B)_{t \geq 0}$ -martingale.

(2) Let  $a \in \mathbb{R}$  and define  $T_a = \inf\{t \geq 0, B_t = a\}$ . Prove that  $\mathbb{E}[e^{-xT_a}] = e^{-a\sqrt{2x}}$  for any  $x \geq 0$ .

(3) Prove that for any  $a > 0$ ,

$$\mathbb{P}(\sup_{0 \leq s \leq t} M_t^\lambda \geq a) \leq \frac{1}{a}. \quad \text{JUST Apply Doob's Inequality}$$

(4) Use the previous result to show the *exponential inequality* : for any  $a > 0$ ,

$$\mathbb{P}(\sup_{0 \leq s \leq t} B_t \geq at) \leq e^{-a^2 t/2}.$$



# Worksheet 4

## Exercise 1

$(X, \mathcal{F}_T)$

$(E[X | \mathcal{F}_t])_{t \geq 0}$  is a  $\mathcal{F}_T$  Martingale?

$$\begin{aligned} * \quad E[|E[X | \mathcal{F}_t]|] &\leq E\left[\underbrace{E[|X| | \mathcal{F}_t]}_{\text{since } |X| \text{ is a martingale}}\right] < \infty \\ &= E[|X|] < \infty \quad \text{finite expectation} \end{aligned}$$

\*  $E[X | \mathcal{F}_t]$  is adapted ( $\mathcal{F}_t$  measurable) because  $X$  is adapted

\* Call  $Y_t = E[X | \mathcal{F}_t]$

$$\text{* } E[Y_{t+s} | \mathcal{F}_t] = E\left[E[X | \mathcal{F}_{t+s}] \mid \mathcal{F}_t\right] \stackrel{\text{tower}}{=} E[X | \mathcal{F}_t] = Y_t$$

$E[Y_{t+s} | \mathcal{F}_t] = Y_t$  Martingale.

## Exercise 2. Ind increments

$X_t$ ,  $X_t - X_s \perp \sigma(X_u, u \leq s)$  and  $E[X_t] = E[X_0]$

$X_t$  integrable.

Any Integrable, Ind Increment, centered Process is Martingale

~~•  $X_t$  is adapted~~

~~•  $E[|X_t|] < \infty$~~

•  $X_t$  is Adapted ( $F_t$  measurable) Natural Filtration

$s > t$

$$E[X_s | F_t] = E[(X_s - X_t) + X_t | F_t]$$

$$= E[X_s - X_t | F_t] + E[X_t | F_t]$$

independent increments +  $X_t$

$X_s - X_t$  ind  $F_t = \sigma(X_u)$

$$= E[X_s - X_t] + X_t$$

$$= E[X_s] - E[X_t] + X_t$$

$$= E[X_0] - E[X_0] + X_t$$

centered.  
$$= X_t$$

## Exercise 2 part b

$$\bullet \mathbb{E} |X_t^2 - \mathbb{E} X_t^2| \leq \mathbb{E} |X_t^2| + \mathbb{E} |\mathbb{E} X_t^2| \leq \mathbb{E} |X_t^2| + \mathbb{E} |X_t^2| < \infty$$

$$\bullet \mathbb{E} [X_t^2 - \mathbb{E} X_t^2 | \mathcal{F}_s]$$

$$= \mathbb{E} (X_t - X_s)^2 | \mathcal{F}_s - \mathbb{E} [ \mathbb{E} (X_t - X_s)^2 | \mathcal{F}_s ]$$

$$= \mathbb{E} [(X_t - X_s)^2 + X_s^2 + 2(X_t - X_s)X_s | \mathcal{F}_s] - \mathbb{E} (X_t - X_s + X_s)^2$$

$$\text{If } X_t - X_s \perp \mathcal{F}_s \Rightarrow (X_t - X_s)^2 \perp \mathcal{F}_s$$

$X_s$  is cte wrt  $\mathcal{F}_s$

$$\text{and } \mathbb{E} X_t - X_s = 0$$

$$= \underline{\mathbb{E} (X_t - X_s)^2} + X_s^2 + 0 - \left\{ \underline{\mathbb{E} (X_t - X_s)^2} + \mathbb{E} X_s^2 + 2 \mathbb{E} (X_t - X_s) X_s \right\}_{=0}$$

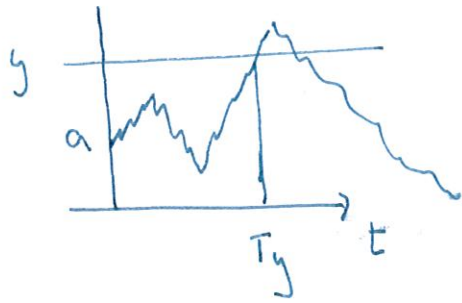
$$= X_s^2 - \mathbb{E} X_s^2 \quad \text{Martingale.}$$



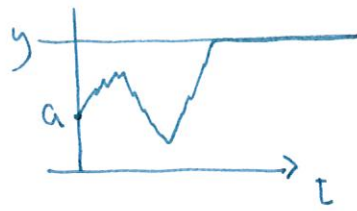


3)  $M_t$

$M_{t \wedge T_y}$  is a UI Martingale.



$T_y$  stopping Time.



$t \wedge T_y$  is stopping Time

$$\begin{aligned} \mathbb{E} M_{t \wedge T_y} &= \mathbb{E} M_t \mathbb{1}_{t < T_y} \mathbb{1}_{T_y < \infty} \\ &+ \mathbb{E} M_t \mathbb{1}_{t > T_y} \mathbb{1}_{T_y < \infty} \\ &+ \mathbb{E} M_t \mathbb{1}_{t < T_y} \mathbb{1}_{T_y = \infty} \end{aligned}$$

$$\begin{aligned} \lim_{t \rightarrow \infty} M_t &= 0 \\ \lim_{t \rightarrow \infty} \mathbb{E} M_{t \wedge T_y} &= 0 + \mathbb{E} \mathbb{1}_{T_y < \infty} y + 0 \\ &= y \mathbb{P}(T_y < \infty) \end{aligned}$$

~~by Doob~~

by Doob

$$\mathbb{E} M_{t \wedge T_y} = \mathbb{E} M_0 = a$$

~~or by optional~~

martingale

$$a = y \mathbb{P}(T_y < \infty)$$

$$\mathbb{P}(T_y < \infty) = \frac{y}{a}$$



## Exercise 4

Product of <sup>ind</sup> BM is Martingale

$$X = B_1 B_2$$

$$(F_t^{(B_1, B_2)})_{t \geq 0} = \sigma(B_1(s), B_2(s) \quad s \leq t)$$

$$X(t) = B_1(t) B_2(t)$$

$$\ast \quad E[|X_T|] = E[|B_1 B_2|] \stackrel{\text{ind}}{=} E[|B_1|] E[|B_2|] < \infty$$

$\ast$   $X_T$  is  $F^{(B_1, B_2)}$  Measurable.

$$\begin{aligned} \circ \quad E[X_t | F_s] &= E[B_1(t) B_2(t) | F_s] \\ &= E[(B_1(t) - B_1(s) + B_1(s))(B_2(t) - B_2(s) + B_2(s)) | F_s] \\ &= E[(B_1(t) - B_1(s))(B_2(t) - B_2(s)) | F_s] + E[(B_1(t) - B_1(s)) B_2(s) | F_s] \\ &\quad + E[B_1(s)(B_2(t) - B_2(s)) | F_s] + E[B_1(s) B_2(s) | F_s] \end{aligned}$$

$$\begin{aligned} &\stackrel{\text{ind incr}}{=} E[B_1(t) - B_1(s) | F_s] E[B_2(t) - B_2(s) | F_s] \\ &\quad + \stackrel{\text{ind incr}}{B_2} E[B_1(t) - B_1(s) | F_s] + B_1(s) E[B_2(t) - B_2(s) | F_s] \\ &\quad + B_1(s) B_2(s) \end{aligned}$$

$$= B_1(s) B_2(s)$$

$$E[B_1(t) B_2(t) | F_s] = B_1(s) B_2(s)$$



# Examples

## 2) Sum of iid centered, integrable

$X_0, X_1, \dots, X_n$  centered iid integrables ( $L^1$ )

$$F_n = \sigma(X_0, X_1, \dots, X_n)$$

$$S_n = \sum_i X_i$$

Triangular or Minkowski

$$\bullet E[|S_n|] = E\left[\left|\sum_i X_i\right|\right] \leq E\left[\sum_i |X_i|\right] = \sum_i E|X_i| < \infty$$

$$|a+b| \leq |a| + |b|$$

$$\|a+b\|_{L^p} \leq \|a\|_{L^p} + \|b\|_{L^p}$$

$$\bullet E[S_{n+1} | F_n] = S_n$$

$$= E[S_n + X_{n+1} | F_n] = E[S_n | F_n] + E[X_{n+1} | F_n]$$

$$= S_n + E[X_{n+1}]$$

centered

$$= S_n \quad \checkmark$$

## Example 2

Variance of sum (ind centered integrable)

$$M_n = S_n^2 - n\sigma^2$$

$$E[M_n] = E[S_n^2] - E[n\sigma^2]$$

~~var~~

$$E[|M_n|] = E[|S_n^2 - n\sigma^2|] \leq E[S_n^2] + E[n\sigma^2] =$$

$$= E\left[\sum_{i \neq j} X_i X_j\right] + n\sigma^2$$

$$\stackrel{\text{ind}}{=} E\left[\sum_{i \neq j} X_i X_j + \sum_i X_i^2\right] + n\sigma^2$$

$$= E\left[\sum_{i \neq j} X_i X_j\right] + E\left[\sum_i X_i^2\right] + n\sigma^2$$

Minkowski

$$\leq E\sum_{i \neq j} (X_i X_j) + E\sum_i (X_i^2) + n\sigma^2$$

$$\stackrel{\text{ind}}{=} \sum_{i \neq j} E[X_i]E[X_j] + n\sigma^2 + n\sigma^2$$

2nd Method.

$$= E S_n^2 + n\sigma^2$$

$$= \text{Var}\left(\sum_i X_i\right) + n\sigma^2$$

ind

$$= \sum_i \text{Var}(X_i) + n\sigma^2$$

$$= n\sigma^2 + n\sigma^2 = 2n\sigma^2$$

$< \infty$

$$E[M_{n+1} | F_n] = E[S_{n+1}^2 - (n+1)\sigma^2 | F_n]$$

$$= E[(S_n + X_{n+1})^2 - (n+1)\sigma^2 | F_n]$$

$$= E[S_n^2 + 2S_n X_{n+1} + X_{n+1}^2 - n\sigma^2 - \sigma^2 | F_n]$$

~~ind~~

$$= -n\sigma^2 - \sigma^2 + 2 \overbrace{E[S_n | F_n]}^{S_n} \overbrace{E[X_{n+1} | F_n]}^{E[X_{n+1}] = 0} + E[S_n^2 | F_n] + E[X_{n+1}^2 | F_n]$$

$$E[X_{n+1}^2] = \sigma^2$$

$$= -n\sigma^2 + E[S_n^2 | F_n]$$

$$= S_n^2 - n\sigma^2 \quad \checkmark$$

### Ex 1.5.1

Any  $f(t, B_t)$  is adapted  $F_t$   
 $B_t^3 - t$

### Example 3 Wald's Martingale.

$X_i$  Sum iid centered integrable

$$M_n = \frac{\exp(\lambda S_n)}{(E[\exp \lambda X_1])^n}$$

$$\bullet E[M_n] = E\left[\frac{\exp \lambda S_n}{(E[\exp \lambda X_1])^n}\right] = E\left[\frac{\exp(\lambda X_1) \exp(\lambda X_2) \dots}{(E[\exp \lambda X_1])^n}\right]$$

$$\text{and} \\ = \frac{E[e^{\lambda X_1}] E[e^{\lambda X_2}] \dots}{(E[e^{\lambda X_1}])^n}$$

$\exp X > 0$

$$\text{iid} \\ = \frac{(E[e^{\lambda X_1}])^n}{(E[e^{\lambda X_1}])^n} = 1.$$

$$\bullet E[M_{n+1} | F_n] = E\left[\frac{e^{\lambda S_{n+1}}}{(E[\exp \lambda X_1])^{n+1}} \middle| F_n\right] = E\left[\frac{e^{\lambda S_n} e^{\lambda X_{n+1}}}{(E[e^{\lambda X_1}])^{n+1}} \middle| F_n\right]$$

$$\text{ind} \\ = \frac{E[e^{\lambda X_{n+1}}]}{E[e^{\lambda X_1}]} \frac{e^{\lambda S_n}}{(E[e^{\lambda X_1}])^n}$$

$$\text{iid} \\ = \frac{e^{\lambda S_n}}{E[e^{\lambda X_1}]^n} = M_n.$$

Example 4: Doob's Martingale  
Accumulating Data

$$M_n = E[X | F_n]$$

$$E[|M_n|] = E[|E[X | F_n]|] \leq E[E[|X| | F_n]] \quad 4/1999$$

11/12/17

Rule 2

$$= E[|X|] < \infty$$

$X \in L_1$

Rule 2:  $E[X] = E[E(X|F)]$

$$E[M_{n+1} | F_n] = E[E[X | F_{n+1}] | F_n] \stackrel{\text{tower}}{=} E[X | F_n] = M_n$$

Proposition

$M$  a Martingale,  $f$  convex  $\Rightarrow f(M)$  is Sub Martingale.

$$M_n = E[M_{n+1} | F_n]$$

$$f(M_n) = f[E[M_{n+1} | F_n]] \leq E[f(M_{n+1}) | F_n]$$

$$f(M_n) \leq E[f(M_{n+1}) | F_n]$$

sub Martingale



iii)

Doob's Inequality

$Y_t$  a non negative submartingale

$$\bullet \mathbb{P}\left(\sup_{0 \leq t \leq T} Y_t \geq c\right) \leq \frac{E(Y_T^p)}{c^p} \quad \begin{array}{l} c > 0 \\ p \geq 1 \end{array}$$

$M_t$  is non negative martingale

Any martingale is also a sub/super martingale

$$\Rightarrow \mathbb{P}\left(\sup_{0 \leq t \leq T} M_t \geq a\right) \leq \frac{E[Y_0]}{a} = \frac{1}{a}$$

$$E Y_0 = e^0 = 1$$

iv)

$$\mathbb{P}\left(\sup B_t > c\right) = \mathbb{P}\left(\sup e^{\lambda B_t} \geq e^{\lambda c}\right)$$

$e^{\lambda B_t}$  is submartingale  
non negative

$B_t$  is martingale

Doob Ineq

$$\leq \frac{E[e^{\lambda B_T}]}{e^{\lambda c}} = e^{\frac{1}{2}\lambda^2 T - \lambda c}$$

optimize right hand side respect to  $\lambda$

$$\frac{d}{d\lambda} \left( \frac{1}{2}\lambda^2 T - \lambda c \right) = \lambda T - c = 0$$

$$\Rightarrow \lambda = \frac{c}{T}$$

$$\Rightarrow \mathbb{P}\left(\sup B_t > c\right) \leq \underbrace{e^{\frac{1}{2} \frac{c^2 T}{T^2} - \frac{c^2}{T}}}_{e^{-\frac{1}{2} \frac{c^2}{T}}}$$

change  $c \rightarrow cT$

$$\Rightarrow \mathbb{P}\left(\sup B_t \geq cT\right) \leq e^{-\frac{1}{2} c^2 T}$$

$$7) \quad M_t = e^{\lambda B_t - \frac{\lambda^2 t}{2}}, \quad \mathbb{E} |M_t| = \mathbb{E} M_t = 1 < \infty$$

$$\mathbb{E} |M_t| \stackrel{CS}{\leq} \left( \mathbb{E} (M_t)^2 \right)^{1/2} = e^{-\frac{\lambda^2 t}{2}} \mathbb{E} e^{\lambda(B_t - B_s + B_s)}$$

$$= e^{-\frac{\lambda^2 t}{2}} \underbrace{\left( \mathbb{E} e^{2\lambda B_t} \right)^{1/2}}_{\text{momentum generator}}$$

$$= e^{-\frac{\lambda^2 t}{2}} \left( e^{\frac{1}{2} 4\lambda^2 t} \right)^{1/2}$$

$$= e^{-\frac{\lambda^2 t}{2} + \lambda^2 t} < \infty \quad \forall t$$

$$\mathbb{E} [M_t | \mathcal{F}_s] = \mathbb{E} \left[ e^{\lambda B_t - \frac{\lambda^2 t}{2}} \mid \mathcal{F}_s \right]$$

$$= \mathbb{E} \left[ e^{\lambda(B_t - B_s + B_s) - \frac{\lambda^2 t}{2}} \mid \mathcal{F}_s \right]$$

$$= e^{\lambda B_s} e^{-\frac{\lambda^2 t}{2}} \underbrace{\mathbb{E} \left[ e^{\lambda(B_t - B_s)} \right]}_{e^{\frac{1}{2} \lambda^2 (t-s)}}$$

$$= e^{\lambda B_s - \frac{\lambda^2 s}{2}} = M_s$$

$M_t$  is UI

because 
$$\sup_t \mathbb{E} |M_t|^p = \sup_t e^{-\frac{\lambda^2 t p}{2}} \underbrace{\mathbb{E} e^{p \lambda B_t}}_{e^{\frac{p^2 \lambda^2 t}{2}}}$$

$$= \sup_t e^{\frac{\lambda^2 t}{2} (p^2 - p)}$$

$< \infty$  just for  $p=1$

so  $\exists M_\infty$  in  $\mathbb{H}^1$

consider  $M_{t \wedge T_a}$  is also UI

$$\mathbb{E} M_{t \wedge T_a} = \mathbb{E} M_{t \wedge T_a} \mathbb{1}_{T_a < \infty} + \mathbb{E} M_{t \wedge T_a} \mathbb{1}_{T_a = \infty}$$

Doob optional

$$1 = \lim_{t \rightarrow \infty} \mathbb{E} M_{t \wedge T_a} \mathbb{1}_{T_a < \infty} + \underbrace{\frac{e^{-\infty}}{0} e^{\lambda B_0}}_0$$

$$\frac{1}{M_{T_a}} = \mathbb{P}(T_a < \infty)$$

$$e^{\frac{\lambda a - \lambda^2 T_a}{2}} =$$

$$1 = \mathbb{E} e^{-\frac{\lambda^2 T_a}{2} + \lambda a}$$

$$e^{-\lambda a} = \mathbb{E} e^{-\frac{\lambda^2 T_a}{2}}$$

$$\frac{\lambda^2}{2} = \kappa \quad \lambda = \sqrt{2\kappa}$$

$$e^{-\sqrt{2\kappa} a} = \mathbb{E} e^{-\kappa T_a}$$



**WORKSHEET 5**  
**WIENER'S INTEGRAL**

In all the exercises,  $(\Omega, \mathcal{A}, \mathbb{P})$  denotes the current probability space and  $(B_t)_{t \geq 0}$  a (real) Brownian motion.

**Exercise 1.**

- (1) Check that the random variable  $Y = \int_0^{+\infty} e^{-s} dB_s$  is well-defined.
- (2) Give the law of  $Y$ .

**Exercise 2.** Find two admissible functions  $f$  and  $g$  such that  $f \leq g$  and

$$\mathbb{P} \left[ \int_0^1 f(s) dB_s > \int_0^1 g(s) dB_s \right] > 0.$$

**Exercise 3.** Let  $f$  be an admissible function. Show that the process  $(\int_0^t f(s) dB_s)_{t \geq 0}$  is a Gaussian process. Compute its mean and its covariance.

**Exercise 4.** Let  $(X_t)_{t \geq 0}$  be given by:

$$\forall t \geq 0, X_t = \int_0^{t^{1/2}} (2s)^{1/2} dB_s.$$

Show that  $(X_t)$  is a Gaussian process. Compute its mean and its covariance. Deduce that  $X$  is a Brownian motion.

**Exercise 5.** Let  $V_0$  be a random variable independent of  $B$  and of Gaussian law  $\mathcal{N}(0, 1/2)$ . We define the process  $(V_t)_{t \geq 0}$  (so-called Ornstein-Uhlenbeck stationary process) by:

$$\forall t \geq 0, V_t = \exp(-t)V_0 + \int_0^t \exp[-(t-s)] dB_s.$$

- (1) Show that  $(V_t)_{t \geq 0}$  is a Gaussian process.
- (2) For any  $a > 0$ , prove that  $(V_{a+t})_{t \geq 0}$  and  $(V_t)_{t \geq 0}$  have the same distribution.

**Exercise 6.** Let  $T > 0$ . Show that

$$\lim_{n \rightarrow +\infty} \mathbb{E} \left[ \left( \sum_{i=1}^n (B_{Ti/n} - B_{T(i-1)/n})^2 - T \right)^2 \right] = 0.$$

**Exercise 7.** Let  $T > 0$ . Show that

$$\int_0^T \left(1 + \frac{B_t}{n}\right)^n dB_t \xrightarrow[n \rightarrow \infty]{L^2} \int_0^T \exp(B_t) dB_t.$$

Check first that the integrals are well-defined.

**Exercise 8.** Let  $T > 0$ . For a given  $n \geq 1$ , we define the process

$$\forall n \geq 0, \forall t \geq 0, B_t^n = \sum_{i=0}^{n-1} B_{T_{i/n}} \mathbf{1}_{(T_{i/n}, T_{(i+1)/n}]}(t).$$

- (1) Prove that  $(B_t^n)_{t \geq 0}$  is a simple process w.r.t. the filtration generated by  $B$ .  
 (2) Show that

$$\lim_{n \rightarrow +\infty} \mathbb{E} \int_0^T |B_t^n - B_t|^2 dt = 0.$$

- (3) What is the limit, in  $L^2(\Omega)$ , of

$$\left( \int_0^T B_t^n dB_t \right)_{n \geq 1} ?$$

- (4) Prove that

$$B_T^2 = 2 \int_0^T B_t^n dB_t + \sum_{i=1}^n (B_{T_{i/n}} - B_{T_{(i-1)/n}})^2$$

- (5) By the previous exercise, deduce that

$$B_T^2 = 2 \int_0^T B_t dB_t + T.$$

## Worksheet 5

$$1) \int_0^t f(s) dB_s \sim N\left(0, \int_0^t f^2(s) ds\right)$$

$$Y = \int_0^\infty e^{-s} dB_s$$

Well define if  $E\left[\int f^2 ds\right] < \infty$

$$= E\left[\int_0^\infty (e^{-s})^2 ds\right] = E\left[\frac{e^{-2s}}{-2}\right]_0^\infty$$

$$= E\left[\frac{1}{2}\right] = \frac{1}{2} < \infty$$

$$\Rightarrow N\left(0, \frac{1}{2}\right)$$

$$2) \mathbb{P}\left[\int_0^1 f(s) dB_s > \int_0^1 g(s) dB_s\right] > 0 \quad \text{s.t. } f < g$$

example  $f = -k$   $g = k$ .

$$\int_0^1 -k dB_s = -k B_1$$

$$\int_0^1 k dB_s = k B_1$$

$$\mathbb{P}\left[-k B_1 > k B_1\right] = \mathbb{P}\left[2k B_1 < 0\right] = \frac{1}{2}$$

### Exercise 4

$$X_t = \int_0^{t^{1/2}} (2s)^{1/2} dB_s \sim N\left(0, \int_0^{\sqrt{t}} 2s ds\right)$$

$$\int_0^{\sqrt{t}} 2s ds = s^2 \Big|_0^{\sqrt{t}} = t \Rightarrow N(0, t)$$

If  $a$  and  $b$  have mean zero

$$\begin{aligned} \text{cov}(a, a+b) &= E[a(a+b)] = E[a^2 + ab] = E[a^2] + E[ab] \\ &= \text{var}(a) + \text{cov}(a, b) \end{aligned}$$

$$\begin{aligned} t < t'; \text{cov}\left(\int_0^t, \int_0^{t'}\right) &= \text{cov}\left(\int_0^t, \int_0^t + \int_t^{t'}\right) \\ &= \text{cov}\left(\int_0^t, \int_0^t\right) + \text{cov}\left(\int_0^t, \int_t^{t'}\right) \\ \text{but } \int_t^{t'} &\perp \int_0^t \\ &= \text{var}\left(\int_0^t\right) \\ &= t \quad \text{because } t < t' \end{aligned}$$

So for  $t' < t$  it will be  $t'$

$$\Rightarrow \text{cov}(X_t, X_{t'}) = \text{mf}(t', t).$$



### Exercise 3

$$\text{Itô} \quad \int_0^t f(s) dB_s = \lim_{n \rightarrow \infty} \sum_{i=1}^n f(t_{i-1}) (B_{t_i} - B_{t_{i-1}})$$

$$S_n = \sum_{i=1}^n \underbrace{f(t_{i-1})}_{\text{constant}} \underbrace{(B_{t_i} - B_{t_{i-1}})}_{N(0, t_i - t_{i-1})} \sim N$$

as  $\Delta_i B$  are independent increments  
and stationary increments

we have a sum of independent Gaussians  
which is also Gaussian

Now

$$M_n = E[S_n] = E\left[\sum \dots\right] = \sum_{i=1}^n f(t_{i-1}) \underbrace{E[\Delta_i B]}_0 = 0$$

$$\sigma_n^2 = E[S_n^2] = E\left[\sum_i \sum_j f(t_{i-1}) f(t_{j-1}) \Delta_i B \Delta_j B\right]$$

$$i \neq j \Rightarrow \Delta_i B \perp \Delta_j B$$

$$= \sum_i \sum_j f(t_{i-1}) f(t_{j-1}) E[\Delta_i B] E[\Delta_j B] = 0$$

$$i=j \Rightarrow E\left[\sum_i f^2(t_{i-1}) (\Delta_i B)^2\right]$$

$$= \sum_i f^2(t_{i-1}) \underbrace{E(\Delta_i B)^2}_{\Delta t_i = \text{var}(\Delta_i B)}$$

$$= \sum_i f^2(t_{i-1}) \Delta t_i$$

conv in distribution

$$M_n \rightarrow M = 0$$

$$\sigma_n^2 \rightarrow \sigma^2 = \int_0^t f^2 ds$$

so the characteristic  
function of  $S_n$   
goes to

$$N\left(0, \int_0^t f^2 ds\right)$$

because mean  $\mathbb{L}^2$   
conv. of the sum  
implies conv  
in Prob.  
and dist.

$$c) \lim \mathbb{E} \left( \sum \left( B_{\frac{Tj}{n}} - B_{\frac{T(j-1)}{n}} \right)^2 - \text{[colorful scribble]} \right)^2$$

$$\lim \mathbb{E} \left[ \left( \sum \left( B_{\frac{Tj}{n}} \right)^2 \right)^2 + T^2 - 2T \sum (B_{\frac{Tj}{n}}) \right]$$

$$+ T^2 - 2T \quad h \frac{T}{n}$$

$$\lim \mathbb{E} \left( \sum \left( \frac{\sqrt{T}}{n} B_1 \right)^2 \right)^2 + T^2 - 2T^2$$

$$\lim \mathbb{E} \left( \frac{T}{n} n B_1^2 \right)^2 + \text{''}$$

$$\lim T^2 \underbrace{\mathbb{E} B_1^4}_{\substack{\text{Kurtosis} \\ \text{normal} \Rightarrow 3}} - 2T^2 = T^2$$

## Exercise 7.

### Dominated Convergence

Let  $\lim f_n = f$  pointwise

s.t.  $|f_n| \leq g(x)$

Then  $\lim_{n \rightarrow \infty} \int |f_n - f| = 0$

also  $\lim \int f_n = \int \lim f_n = \int f$

---

We will show that  $\lim \left(1 + \frac{B_t}{n}\right)^n = e^{B_t}$

That in  $\left( \int_0^t \left(1 + \frac{B_t}{n}\right)^n dB_t - \int_0^t e^{B_t} dB_t \right)^2$  both are well defined

means  $f_n = \left(1 + \frac{B_t}{n}\right)^n - e^{B_t}$

p.w.  $\lim f_n = 0$

as  $(a-b)^2 \geq 0$   
 $a^2 + b^2 - 2ab \geq 0$

$a^2 + b^2 \geq 2ab$   
 $2a^2 + 2b^2 \geq (a+b)^2 \geq (a-b)^2$

We can bound  $f_n$  by  $2 \left(1 + \frac{B_t}{n}\right)^{2n} + 2 e^{2B_t} = 4e^{2B_t}$  because well defined

So we can say  $\lim \int f_n = \int \lim f_n = \int f$ .

$$\lim \int \left| \left(1 + \frac{B_t}{n}\right)^n - e^{B_t} \right|^2 dB_t$$

$$= \int 0 dB_t = 0.$$

For checking that integrals are well defined

$$\int_0^t E[C_s]^2 dt < \infty$$

$$E[e^{\lambda z}] = e^{\mu\lambda + \frac{1}{2}\sigma^2\lambda^2}$$

$$E[e^{i\lambda z}] = e^{i\mu\lambda - \frac{1}{2}\sigma^2\lambda^2}$$

$$C_s = e^{Bt} \quad \int_0^t \underbrace{E[e^{2Bt}]}_{\substack{\text{Momentum} \\ \text{Generating} \\ \text{function}}} dt = \int_0^t e^{\frac{1}{2}(t)^2(2)^2} dt =$$

$$= \int_0^t e^{2t} dt < \infty$$

$$C_s = \left(1 + \frac{Bt}{n}\right)^{2n} = e^{2n \log\left(1 + \frac{Bt}{n}\right)} \quad \text{as } 1+x \leq e^x$$
$$\leq e^{2n \frac{Bt}{n}} = e^{2Bt} \quad \text{(same as before)} \quad \log(1+x) \leq x$$

so both integrals are well defined.

## Exercise 5

$$V_0 \sim N(0, \frac{1}{2})$$

$$V_t = \underbrace{e^{-t} V_0}_{N(0, \frac{e^{-t}}{2})} + \underbrace{e^{-t} \int_0^t e^s dB_s}_{N(0, e^{2t} \int_0^t e^{2s} dB_s)}$$

sum of normal is normal

$$\text{if ind} \Rightarrow N(m_1 + m_2, \text{var}_1 + \text{var}_2)$$

$$\text{not ind} \Rightarrow N(m_1 + m_2, \text{var}_1 + \text{var}_2 + 2 \text{cov}_{12})$$

$$\Rightarrow V_t \text{ is normal}$$

$V_{att}$

$$E V_{att} = 0 + 0 = 0$$

$$\text{var } V_{att} = e^{-2(t+a)} \text{var } V_0 + e^{-2(t+a)} \text{var} \int_0^{(t+a)} e^s dB_s$$

$$= e^{-2(t+a)} \frac{1}{2} + e^{-2(t+a)} \frac{1}{2} \int_0^{t+a} e^{2s} ds$$

$$= e^{-2(t+a)} \frac{1}{2} + e^{-2(t+a)} \frac{1}{2} (e^{2(t+a)} - 1)$$

$$= \frac{1}{2} \quad \text{ind } t, a, \dots$$

$$\text{cov}(V_t, V_{t'}) = e^{-t-t'} \frac{1}{2} + e^{-t-t'} \left( \frac{e^{2t} - 1}{2} \right)$$

$$\text{cov}(V_t, V_{t'})$$

$$\text{cov}(A+B, C+D) = \text{cov} A C + \text{cov} A D + \text{cov} B C + \text{cov} B D$$

$$\text{cov}(e^{-t} V_0, e^{-t'} V_0) = e^{-t-t'} \frac{1}{2}$$

$$\text{cov}(e^{-t} V_0, e^{-t'} \int_0^{t'} W_s ds) = e^{-2t'} \cdot 0 \quad V_0 \perp \int_0^t$$

$$\text{cov}(e^{-t} \int_0^t W_s ds, e^{-t'} V_0) = 0$$

$$\begin{aligned} \text{cov}(e^{-t} \int_0^t W_s ds, e^{-t'} \int_0^{t'} W_s ds) &= e^{-t-t'} \left( \text{cov}\left(\int_0^t, \int_0^t\right) + \underbrace{\text{cov}\left(\int_0^t, \int_0^{t'}\right)}_{\substack{0 \\ \text{ind incr.}}} \right) \\ &= e^{-t-t'} \left( \text{var} \int_0^t \right) \end{aligned}$$

$$\Rightarrow \text{cov} V_t, V_{t'} = e^{-t-t'} \frac{1}{2} + e^{-t-t'} \int_0^{t'} e^{2s} ds$$

$$= e^{-t-t'} \frac{1}{2} + e^{-t-t'} \left( \frac{e^{2t'} - 1}{2} \right)$$

$$= e^{-t-t'} \left( \frac{e^{2t'}}{2} \right)$$

$$\text{if } t' = t \Rightarrow \text{cov} V_t, V_t = \text{var} V_t = \frac{1}{2}$$

match with

$$\text{var} V_t =$$

## WORKSHEET 6

### ITO'S FORMULA

In all the exercises,  $(\Omega, \mathcal{A}, \mathbb{P})$  denotes the current probability space and  $(B_t)_{t \geq 0}$  a (real) Brownian motion.

**Exercise 1.** For  $\lambda$  and  $\theta$  in  $\mathbb{R}$ , we consider the process

$$\forall t \geq 0, X_t = \exp(-\lambda t) \cos(\theta B_t).$$

- (1) Compute  $dX_t$  for  $t \geq 0$ .
- (2) What are the values of  $(\lambda, \theta)$  for which the  $dt$ -term in  $dX_t$  vanishes?
- (3) Deduce  $\mathbb{E}[\cos(\theta B_t)]$  for  $t \geq 0$ .

**Exercise 2.** For  $r$  and  $\sigma$  in  $\mathbb{R}$ , we consider the process

$$\forall t \geq 0, X_t = \exp(rt + \sigma B_t).$$

- (1) Compute  $dX_t$  for  $t \geq 0$ .
- (2) What are the values of  $(r, \sigma)$  for which the  $dt$ -term vanishes?
- (3) For the values of  $r$  and  $\sigma$  obtained above, show that, for all  $0 \leq s < t$ ,

$$\mathbb{E}[X_t | \mathcal{F}_s] = X_s,$$

where  $\mathcal{F}_s$  is the  $\sigma$ -field generated by  $(B_u)_{0 \leq u \leq s}$ .

**Exercise 3.** Let  $n$  be an integer larger than 1.

- (1) Show that

$$\forall t \geq 0, B_t^{2n} = 2n \int_0^t B_s^{2n-1} dB_s + n(2n-1) \int_0^t B_s^{2n-2} ds.$$

- (2) Deduce that

$$\mathbb{E}(B_1^{2n}) = (2n-1)\mathbb{E}(B_1^{2n-2}).$$

- (3) Let  $Z$  be an  $\mathcal{N}(0, 1)$  Gaussian variable. Deduce from the above expression that

$$\mathbb{E}(Z^{2n}) = [(2n)!] / [2^n \times n!].$$

**Exercise 4.** Show that the following processes are martingales w.r.t. the filtration generated by  $B$ :

- (1)  $\forall t \geq 0, X_t = \exp(t/2) \cos(B_t)$ .
- (2)  $\forall t \geq 0, Y_t = \exp(t/2) \sin(B_t)$ .
- (3)  $\forall t \geq 0, Z_t = (B_t + t) \exp(-B_t - t/2)$ .
- (4)  $\forall t \geq 0, W_t = B_t^3 - 3tB_t$ .

**Exercise 5.** Let  $(B_t)_{t \geq 0}$  be an  $(\mathcal{F}_t)_{t \geq 0}$ -Brownian motion. Show that  $(B_t^4 - 6tB_t^2 + 3t^2)_{t \geq 0}$  is a martingale w.r.t. to the filtration  $(\sigma(B_s, s \leq t))_{t \geq 0}$ .

**Exercise 6.** Let  $(B_t)_{t \geq 0}$  be an  $(\mathcal{F}_t)_{t \geq 0}$ -Brownian motion and  $(b_t)_{t \geq 0}$  be a continuous and  $(\mathcal{F}_t)_{t \geq 0}$ -adapted process. Set

$$\forall t \geq 0, X_t = \int_0^t b_s ds + B_t.$$

We assume that there exist two constants  $K$  and  $\lambda$  such that

$$\forall t \geq 0, \forall \omega \in \Omega, |b_t(\omega)| \leq K, b_t(\omega)X_t(\omega) \leq -(\lambda/2)X_t^2(\omega).$$

- (1) Show that for all  $T \geq 0$ ,  $\sup_{0 \leq t \leq T} \mathbb{E}[X_t^2] < +\infty$ .
- (2) Applying Itô's formula to  $(\exp(\lambda t)X_t^2)_{t \geq 0}$ , show

$$\sup_{t \geq 0} \mathbb{E}[X_t^2] < +\infty.$$



# Worksheet 6

Exercise 1.  $X_t = e^{-\lambda t} \cos(\theta B_t)$

a)  $df(t, x) = f_1 dt + f_2 dx + \frac{1}{2} [ \overset{\text{order 2}}{f_{11} dt^2} + f_{22} dx^2 + 2 f_{12} dt dx ]$  order 2

$f(t, B_t) = X_t$       $f(t, x) = e^{-\lambda t} \cos \theta x$

$f_1 = -\lambda X_t$       $f_2 = -e^{\lambda t} \theta \sin \theta x$   ~~$f_1 = -\lambda X_t$~~

$f_{22} = -\theta^2 \cancel{e^{\lambda t}} e^{\lambda t} \cos \theta x$

plug in  $x = B_t$   
keep  $dB_t^2 = dt$

$$df(t, B_t) = dX_t = \left[ -\lambda X_t - \frac{\theta^2}{2} X_t \right] dt$$

$$+ -e^{\lambda t} \theta \sin \theta B_t dB_t$$

$$dX_t = \left( -\lambda - \frac{\theta^2}{2} \right) X_t dt - e^{\lambda t} \theta \sin \theta B_t dB_t$$

b)  $-\lambda - \frac{\theta^2}{2} = 0$  ;  $\lambda = -\frac{\theta^2}{2}$

c) It is a Martingale  $X_t = X_0 + \int_0^t -\theta e^{\lambda t} \sin \theta B_t dB_t$

$$E[X_t | F_0] = X_0 = e^{-\lambda t} \cos \theta B_t$$

$$E[E[X_t | F_0]] = E[X_t] = E[X_0]$$

$$X_0 = e^0 \cos(\theta B_0) = 1$$

$$1 = E[e^{-\lambda t} \cos(\theta B_t)]$$

$$e^{\lambda t} = E[\cos(\theta B_t)]$$

## Exercise 2

a)  $X_t = e^{rt + \sigma B_t}$

$$df(t, x) = f_1 dt + f_2 dx +$$

$$X_t = f(t, B_t)$$

$$\frac{1}{2} [ f_{11} dt^2 + f_{22} dx^2 + 2 f_{12} dt dx ]$$

$$f(t, x) = e^{rt + \sigma x}$$

$$f_1 = rf \quad f_2 = \sigma f \quad f_{22} = \sigma^2 f$$

$$df(t, B_t) = rf dt + \sigma f dB_t + \frac{1}{2} \sigma^2 f dt$$

$$= \left( r + \frac{1}{2} \sigma^2 \right) f dt + \sigma f dB_t$$

$$dX_t = \left( r + \frac{1}{2} \sigma^2 \right) X_t dt + \sigma X_t dB_t.$$

b)  $r = -\frac{\sigma^2}{2}$

c)  $X_t = X_0 + \int_0^t \sigma X_s dB_s$  is Martingale.

$\forall t$   $X_t$  is  $\mathcal{F}_t$  measurable.

$$E[X_t | \mathcal{F}_s] = X_0 + E\left[\int_0^t \sigma X_s dB_s \mid \mathcal{F}_s\right]$$

$$= X_0 + \int_0^s \sigma X_s dB_s + E\left[\int_s^t \sigma X_s dB_s \mid \mathcal{F}_s\right]$$

$$= X_0 + \dots + E\left[\int_s^t \sigma X_s dB_s\right]$$

$$\underbrace{\int_s^t E[\sigma^2 X_s^2] ds}_{= 0}$$

$$= X_0 + \int_0^t \sigma X_s dB_s.$$

### Exercise 3

$$a) f(t, x) = x^{2n} \quad f_1 = 0 \quad f_2 = 2n x^{2n-1} \quad f_{22} = 2n(2n-1)x^{2n-2}$$

$$df(t, B_t) = \left( f_1 + \frac{1}{2} f_{22} \right) dt + f_2 dB_t.$$

$$B_t^{2n} - B_0^{2n} = \int_0^t \frac{1}{2} 2n(2n-1) B_t^{2n-2} dt + \int_0^t 2n B_t^{2n-1} dB_t$$

$$b) E[B_t^{2n}] = n(2n-1) E \int_0^t B_t^{2n-2} dt + \underbrace{2n E \int_0^t B_t^{2n-1} dB_t}_{=0}$$

scaling  $c^{1/2} B_t \stackrel{\text{L}}{=} B_{ct}$

$$(t^{1/2} B_1)^{2n-2} \stackrel{\text{L}}{=} B_t^{2n-2}$$

$$t^{n-1} B_1^{2n-2} \stackrel{\text{L}}{=} B_t^{2n-2}$$

$$= n(2n-1) E[B_1^{2n-2}] \int_0^t t^{n-1} dt$$

$$\underbrace{E[B_t^{2n}]}_{=} = n(2n-1) E[B_1^{2n-2}] \frac{t^n}{n}$$

$$E[(t^{1/2} B_1)^{2n}]$$

$$\cancel{t^n} E[B_1^{2n}] = (2n-1) E[B_1^{2n-2}] \cancel{t^n}$$

$$E[B_1^{2n}] = (2n-1) E[B_1^{2n-2}] //$$

c) induction

$$E[Z^0] = \frac{0!}{2^0 0!} = 1.$$

previous result.  $E[Z^{2n}] = (2n-1) E[Z^{2n-2}]$

suppose it's true for  $n-1$   $= (2n-1) \frac{2^{n-1} (n-1)!}{2^{n-1} (n-1)!}$

$$= \frac{2n(2n-1)(2n-2)!}{2 \cdot 2^{n-1} n(n-1)!}$$

$$= \frac{2n!}{2^n n!}$$

### Exercise 4

$$\text{iii) } X_t = (B_t + t) e^{-B_t - t/2} = f(t, B_t)$$

$$f(t, x) = (x + t) e^{-x - t/2}$$

$$\int_0^t df(t, x)_{B_t} = \int_0^t (f_1 + \frac{1}{2} f_{22}) dt + \int_0^t f_2 dB_t$$

$$f(0, B_0) = 0 \quad f_1 = e^{-x - t/2} - \frac{1}{2} (x + t) e^{-x - t/2}$$

$$f_2 = e^{-x - t/2} - (x + t) e^{-x - t/2}$$

$$f_{22} = (x + t) e^{-x - t/2} - 2 e^{-x - t/2}$$

$$X_t = 0 + \int_0^t 0 dt + \underbrace{\int_0^t e^{-B_t - t/2} (1 - (B_t + t)) dB_t}_{\text{M\ddot{o}}}_0$$

$\text{M\ddot{o}}$  is martingale.

Example 3

$$K = x^2 + 2x + 1$$

$$K = (x+1)^2$$

Let  $x = 1$  then  $K = (1+1)^2 = 2^2 = 4$

Let  $x = 2$  then  $K = (2+1)^2 = 3^2 = 9$

Let  $x = 3$  then  $K = (3+1)^2 = 4^2 = 16$

Let  $x = 4$  then  $K = (4+1)^2 = 5^2 = 25$

Let  $x = 5$  then  $K = (5+1)^2 = 6^2 = 36$

Let  $x = 6$  then  $K = (6+1)^2 = 7^2 = 49$

Let  $x = 7$  then  $K = (7+1)^2 = 8^2 = 64$

Let  $x = 8$  then  $K = (8+1)^2 = 9^2 = 81$

Let  $x = 9$  then  $K = (9+1)^2 = 10^2 = 100$

Let  $x = 10$  then  $K = (10+1)^2 = 11^2 = 121$

Let  $x = 11$  then  $K = (11+1)^2 = 12^2 = 144$

Let  $x = 12$  then  $K = (12+1)^2 = 13^2 = 169$

Let  $x = 13$  then  $K = (13+1)^2 = 14^2 = 196$

Let  $x = 14$  then  $K = (14+1)^2 = 15^2 = 225$

Let  $x = 15$  then  $K = (15+1)^2 = 16^2 = 256$

### Exercise 5

$$X_t = B_t^4 - 6t B_t^2 + 3t^2$$

$$f(t, x) = x^4 - 6t x^2 + 3t^2$$

$$f_1 = -6x^2 + 6t$$

$$f_2 = 4x^3 - 12tx$$

$$f_{22} = 12x^2 - 12t$$

$$X_t - X_0 = \int_0^t -6B_s^2 + 6s + \frac{1}{2} (12B_s^2 - 12s) dt \\ + \int_0^t (4B_s^3 - 12sB_s) dB_s$$

$$= \int_0^t c dt + \underbrace{\int_0^t (4B_s^3 - 12sB_s) dB_s}_{\text{IT\hat{o} is Martingale.}}$$

### Exercise 6

$$X_t = \int_0^t b_s ds + B_t$$

$$X_t^2 = \left( \int_0^t b_s ds \right)^2 + B_t^2 + 2B_t \int_0^t b_s ds \quad |b_t| \leq K$$

$$\left| \int_0^t b_s ds \right| \leq \int_0^t |b_s| ds \leq Kt$$

$$\left| \int_0^t b_s ds \right|^2 \leq K^2 t^2$$

$$\text{so, } E \left[ \int_0^t b_s ds \right]^2 \leq K^2 t^2$$

$$E [B_t^2] = t$$

$$E \left[ B_t \int_0^t b_s ds \right] = \int_{-\infty}^{\infty} B_t \cdot \int_0^t b_s ds \, dP = \langle B_t, \int_0^t b_s ds \rangle$$

$$\stackrel{\text{CS}}{\leq} (E B_t^2)^{1/2} (E (\int_0^t b_s ds)^2)^{1/2}$$

$$= t^{1/2} (K^2 t^2)^{1/2}$$

$$E[X_t^2] = K^2 t^2 + t + t^{1/2} K t$$

$$\sup_t E[X_t^2] = K^2 T^2 + T + T^{3/2} K$$

$$\text{ii) } X_t = \int_0^t b_s ds + B_t \quad f(t, x) = e^{\lambda t} x^2$$

$$f_1 = \lambda f \quad f_2 = 2 e^{\lambda t} x \quad f_{22} = 2 e^{\lambda t}$$

$$e^{\lambda t} X_t^2 - e^{\lambda 0} X_0^2 = \int_0^t f_1 d\tau + \int_0^t f_2 dX_\tau + \frac{1}{2} \left[ \dots f_{22} dX_\tau^2 \right]$$

$$= \int_0^t \lambda e^{\lambda \tau} X_\tau^2 d\tau + \int_0^t 2 e^{\lambda \tau} X_\tau dX_\tau + \int_0^t e^{\lambda \tau} dX_\tau^2$$

$$\int_0^t \lambda e^{\lambda \tau} X_\tau^2 d\tau + \int_0^t 2 e^{\lambda \tau} X_\tau dX_\tau + \int_0^t e^{\lambda \tau} dX_\tau^2$$

$$\ast dX_t = b dt + dB_t \quad \ast dX_t^2 = b dt^2 + dB_t^2 + 2b dt dB_t$$

$$e^{\lambda t} X_t^2 - e^{\lambda 0} X_0^2 = \lambda e^{\lambda t} X_t^2 dt + 2 e^{\lambda t} X_t [b dt + dB_t]$$

$$+ \frac{1}{2} \left[ \dots + 2 e^{\lambda t} [b dt^2 + dB_t^2 + 2b dt dB_t] \right]$$